

A Model for Retirement Theory and Estimation

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- Methode for analyzing retirement used in latest report from Danish Economics Councils (Link: http://www.dors.dk/graphics/Synkron-Library/Publikationer/Rapporter/For%20E5r_2013/Trykt/F13_Kapitel_3.pdf)
- Collaboration between Danish Economics Councils (DØRS) and DREAM (Working paper: <http://www.dors.dk/graphics/Synkron-Library/Publikationer/Arbejdspapirer/Arbejdspapir%20om%20tilbagebetraktningmodel.pdf>)

- 1 The model
- 2 Estimation
- 3 Analysis
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Stock & Wise (1990)

Income for a person retiring at age r :

$$Y_s \text{ for } s < r$$

$$B_s(r) \text{ for } s \geq r$$

Stock & Wise (1990)

Utility function:

$$U_w(Y_s) = Y_s^\gamma + w_s$$

$$U_r(B_s(r)) = (k B_s(r))^\gamma + \xi_s$$

where

$$w_s = \rho w_{s-1} + \epsilon_{ws}$$

$$\xi_s = \rho \xi_{s-1} + \epsilon_{\xi s}$$

Stock & Wise (1990)

Remarks:

- Utility is a function of income not consumption (= credit rationing or myopic behavior). No consumption smoothing
- Utility function is not general (Y^γ is used in stead of $\frac{1}{\gamma} Y^\gamma$)
- Utility function is stocastic (health costs)
- Parameter k is identical for all

Arnberg & Stephensen (2013)

- Perfect capital market. Consumption smoothing
- General utility function (CRRA)
- Utility function not stochastic (only time of death is random)
- Parameter k is not identical for all (random coefficient)

Income

As S&W: Income for a person retiring at age r :

$$Y_s \text{ for } s < r$$

$$B_s(r) \text{ for } s \geq r$$

Individual $B_s(r)$ are calculated using all the tricks of the trade.

The Consumer: Inspired by S&W

Utility at age 0, given retirement at age r :

$$U_0(r; k) = \phi V_0(r; k) + \epsilon_r$$

$$V_0(r; k) = \sum_{s=1}^T \frac{(\gamma_s(r; k) c_s)^{1-\rho}}{1-\rho} \beta_s$$

$$\gamma_s(r; k) = \begin{cases} 1 & \text{for } s < r \\ k > 1 & \text{for } s \geq r \end{cases}$$

$$\beta_t \equiv \prod_{s=0}^t \frac{1 - \mu_s}{1 + \theta}$$

The Consumer (cont...)

Budget restriction:

$$A_s = (1 + i_s)A_{s-1} + y_s - c_s$$

The interest rate is given by (similar to Yaari (1965) and Blanchard (1985)):

$$(1 + i_s)(1 - \mu_s) \equiv 1 + i$$

Income:

$$y_s = y_s(r) = \begin{cases} Y_s & \text{for } s < r \\ B_s(r) & \text{for } s \geq r \end{cases}$$

Analytical solution

Optimal behavior:

$$U_0(r; k) = \phi \frac{\hat{V}_0(r; k)^{1-\rho}}{1-\rho} + \epsilon_r$$

Indirect utility function:

$$\hat{V}_0(r; k) = \frac{A_0 + H_0(r)}{P_0(r; k)}$$

$$H_0(r) \equiv \sum_{s=1}^T y_s(r) R_s, \quad R_s = \prod_{v=0}^s \frac{1 - \mu_v}{1 + i}$$

$$P_0(r; k) \equiv \left[\sum_{s=1}^T \gamma_s(r; k)^{\frac{1-\rho}{\rho}} \beta_s^{\frac{1}{\rho}} R_s^{\frac{\rho-1}{\rho}} \right]^{\frac{\rho}{\rho-1}}$$

Deterministic case

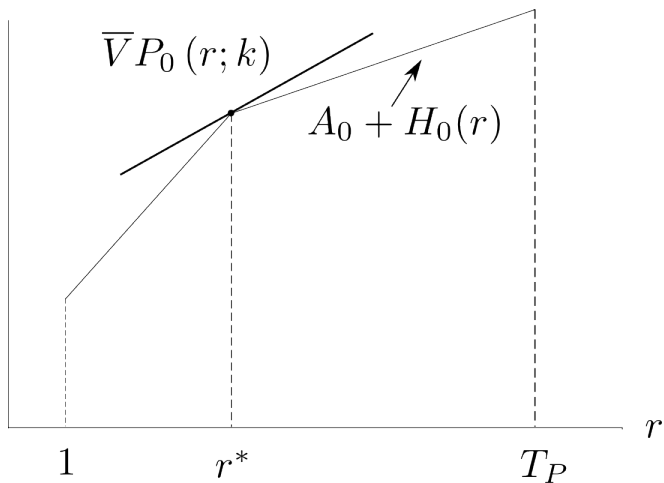
Indifference curve:

$$\hat{V}_0(r; k) = \frac{A_0 + H_0(r)}{P_0(r; k)} = \bar{V}$$

such that:

$$A_0 + H_0(r) = \bar{V}P_0(r; k)$$

Deterministic case: Indifference curves



Estimation: Statistical Model

- Utility based discrete-choice-model with random coefficient:

$$U_0(r; k) = \phi V_0(r; k) + \epsilon_r$$

Could choose many models: logit, probit, mixed logit osv.

- We choose the simplest: logit model with random coefficient
(k) \Rightarrow the ϵ'_r -s are independent extreme value distributed

Logit assumption

The logit assumption implies that:

$$P(r|k, \phi) = \frac{\exp\left(\phi \frac{\hat{V}_0(r;k)^{1-\rho}}{1-\rho}\right)}{\sum_{s=1}^{T_P} \exp\left(\phi \frac{\hat{V}_0(s;k)^{1-\rho}}{1-\rho}\right)}$$

Estimation of random coefficient

- Non-parametric estimation of variation in k .
- Method similar to Kenneth Train (2007) “A Recursive Estimator for Random Coefficient Models”
- The method exploits that:

Expected conditional distribution = distribution

Estimation of random coefficient

Assume a population distribution $p(k)$. This is the distribution we are looking for.

Define for person j :

$$P_j(r_j|k_j) = \pi(r_j|k_j; \phi, x_j)$$

where π is the logit-specification and

$$x_j = \left(A_0^j + H_0^j(1), \dots, A_0^j + H_0^j(T_P) \right)$$

- We know x_j, r_j
- We do not know k_j

Expected Conditional Distribution = Distribution

Conditional distribution: how is k_j distributed given what person j actually did? Bayes rule:

$$P_j(k_j|r_j) = \frac{\pi(r_j|k_j; \phi, x_j) p(k)}{P(r_j; \phi, x_j)}$$

where

$$P(r_j; \phi, x_j) \equiv \int_1^{\infty} \pi(r_j|\kappa; \phi, x_j) p(\kappa) d\kappa$$

or

$$P_j(k_j|r_j) \propto \pi(r_j|k_j; \phi, x_j) p(k)$$

Expected Conditional Distribution = Distribution

You can show that (I tried - and Train did):

$$E [P_j (k|r_j)] = p(k)$$

$$p(k) \Rightarrow P_j (k|r_j) \Rightarrow E [P_j (k|r_j)] \simeq \frac{1}{N} \sum_{j=1}^N P_j (k|r_j) \Rightarrow p(k)$$

This implies that $p(k)$ is a fix point.

Strong intuition: if you take the average of everybodys best guess on the k -distribution, you get the k -distribution.

Fix Point: Successive over-relaxation

How do we find the fix point?:

- Start with some distribution $p^1(k)$
- In iteration i : Calculate for all j : $P_j^i(k|r_j) = \frac{\pi(r_j|k; \phi, x_j) p^i(k)}{P(r_j; \phi, x_j)}$
- In iteration i : Calculate new distribution
 $p^{i+1}(k) = \lambda p^i(k) + (1 - \lambda) \sum_{j=1}^n P_j^i(k|r_j)$
- Continue till convergence

Analysis: Data

Danish register data for the cohorte that had age 59 in 2001

For each individual j we know

- time of retirement r_j
- Income Y_{js} for $s < r_j$
- Pensions $B_s(r_j)$ for $s \geq r_j$

We calculate

- Income Y_{js} for $s \geq r_j$ (using average age-profile)
- Pensions $B_s(r)$ for $r \neq r_j$ (using actualial math)

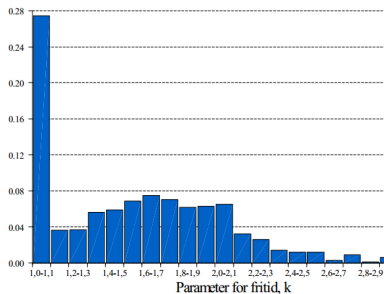
Analysis: The k-distribution

	Estimate
1. quartile	1.35
Median	1.57
3. quartile	1.82

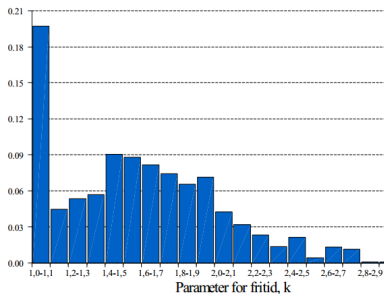
Remark: We assume $\rho = 2$. It is estimated that $\phi = 8$ mio.

Analysis: The k-distribution

a) Mænd



b) Kvinder



Analysis: Retirement (pct.)

Age	Actual	Expected
60	31	30
61	4	3
62	26	19
63	8	6
64	4	4
65	10	7
66	4	12
67+	13	20

Further research

- Peer-group-effect
- Age-effect (k increases as you get older)